Suppose that we want to estimate the consumption function

$$
c_{t}^{*}=\beta_{1}+\beta_{2} y_{i}^{*}
$$

where $c_{t}^{*}$ and $y_{*}^{*}$ are true but unobservable consumption and income, respectively. The observable values $c_{t}$ and $y_{t}$ are subject to errors of measurement and are related to the true values as follows:

$$
c_{t}=c_{t}^{*}+v_{t} \quad y_{t}=y_{t}^{*}+u_{t}
$$

where $v_{t}$ and $u_{t}$ are each independent and identically distributed random variables with $N\left(0, \sigma_{v}^{2}\right)$ and $N\left(0, \sigma_{w}^{2}\right)$ distributions, respectively. Data on $c_{q}$, $y_{t}$, and two potential instrumental variables, $i_{t}$ (investment) and $g_{f}$ (government expenditure), appear in Table 14.3. (KDATA'H ON WERSITE')
(a) Find least squares estimates of $\beta_{1}$ and $\beta_{2}$. (OLS)
(b) Find instrumental variables estimates for $\beta_{1}$ and $\beta_{2}$ by using the following instruments:
(i) $i_{t}$
(ii) g ,
(iii) $x_{i}=i_{i}+g_{1}$
(iv) $i_{t}$ and $g_{t}$

Comment on the alternative estimates and their standard errors.
(c) A number of variations of Hausman's specification test for testing for contemporaneous correlation between $y_{t}$ and the composite error $v_{t}-\beta_{2} u_{t}$ are possible. These variations depend on (i) the instrumental variable (IV) estimator that is used, and (ii) whether the error variance $\operatorname{var}\left(v_{t}-\beta_{2} u_{t}\right)=$ $\sigma^{2}$ is estimated from least squares procedures or instrumental variables procedures. Using estimates for $\beta_{2}$, carry out Hausman's test for the following cases. Comment on the outcome.

| Case | IV Estimator | Choice of $\hat{\sigma}^{\mathbf{2}}$ |
| :---: | :--- | :--- |
| 1 | Uses $i_{t}$ | IV |
| 2 | Uses $g_{t}$ | IV |
| 3 | Uses $x_{t}$ | Least squares |
| 4 | Uses $x_{i}$ | IV |
| 5 | Uses $i_{t}$ and $g_{t}$ | IV |

Table 14.3 Hypothetical Data for $i, g, c$, and $y$

| Observation | $i$ | $g$ | $c$ | $y$ |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 1.5 | 0.5 | 15.30 | 17.30 |
| 2 | 1.4 | 0.6 | 19.91 | 21.91 |
| 3 | 1.5 | 0.7 | 20.94 | 22.96 |
| 4 | 1.4 | 0.8 | 19.66 | 21.86 |
| 5 | 1.5 | 0.9 | 21.32 | 23.72 |
| 6 | 1.4 | 1.0 | 18.33 | 20.73 |
| 7 | 1.6 | 1.0 | 19.59 | 22.19 |
| 8 | 1.5 | 1.1 | 21.30 | 23.90 |
| 9 | 1.6 | 1.2 | 20.93 | 23.73 |
| 10 | 1.6 | 1.2 | 21.64 | 24.44 |
| 11 | 1.7 | 1.3 | 21.90 | 24.90 |
| 12 | 1.6 | 1.4 | 20.50 | 23.50 |
| 13 | 1.8 | 1.4 | 22.83 | 26.05 |
| 14 | 1.7 | 1.5 | 23.49 | 26.69 |
| 15 | 1.9 | 1.5 | 24.20 | 27.60 |
| 16 | 1.8 | 1.6 | 23.05 | 26.45 |
| 17 | 2.0 | 1.6 | 24.01 | 27.61 |
| 18 | 1.9 | 1.7 | 25.83 | 29.43 |
| 19 | 2.0 | 1.8 | 25.15 | 28.95 |
| 20 | 2.0 | 1.8 | 25.06 | 28.86 |

