

EXTENSIONS OF D-F TESTS.

1. TO TEST FOR UMT ROOTS IN THE PRESENCE OF
 - a) SERIALLY CORRELATED ERRORS
 - b) HETERO SKEDASTICITY

⇒ Phillips - Perron test.
2. TO TEST FOR MULTIPLE UNIT ROOTS:

if you test for one root, while in reality there are 2 or more, the empirical size of test > nominal.
that is not 5% but 20%.

therefore better chance to find all roots if you first test for one, then for 2....

Dickey - Pantula: sequential test has
preserves the nominal size.

(sequence of t-test has more power than a joint F-test)

1. $H_0 : m$ roots against $H_A : (m-1)$ roots
if rejected, then

2. $H_0 : (m-1)$ roots against $(m-2)$
etc.

STRUCTURAL CHANGE (REGIME SWITCH)

low power of unit root test: cannot distinguish between a unit root and a regime switch.

true: $y_t = 0.5 y_{t-1} + \varepsilon_t + D_L$

$$D_L = \begin{cases} 0 & \text{for } t = 1 \dots 50 \\ 3 & \text{for } t = 51 \dots 100 \end{cases}$$

fitted: $y_t = a_0 + a_1 y_{t-1} + \ell_t$

\hat{a}_1 is biased to 1!

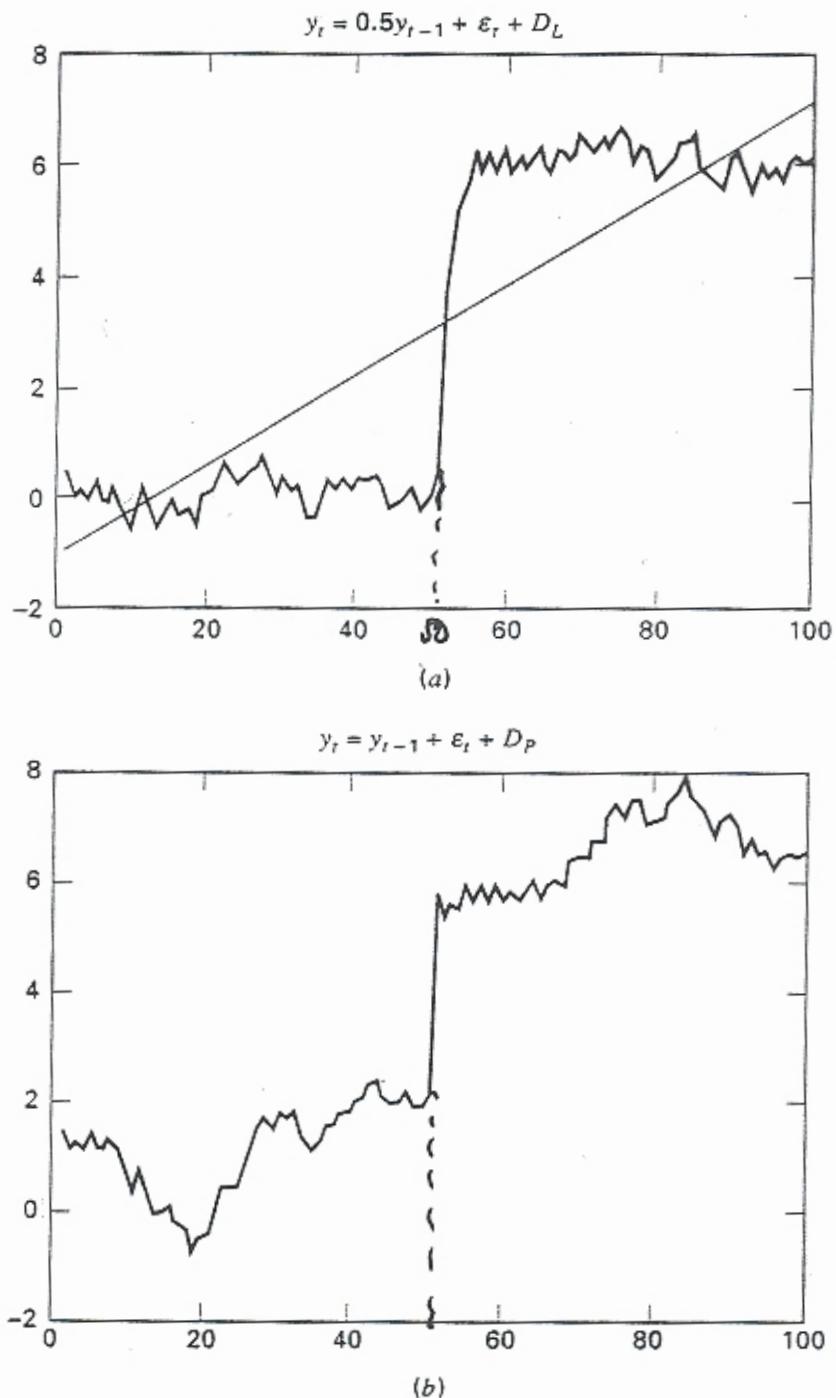
test is biased towards accepting H_0 : unit root even though the series is stationary.

The fitted model mimics the best OLS regression

$$y_t = a_0 + a_2 t + \varepsilon_t$$

$$\Leftrightarrow y_t = y_0 + a_0 t + \sum_{i=1}^t \varepsilon_i$$

Figure 4.4 Two models of structural change.



tures the property that "low" values of y_t (i.e., those fluctuating around zero) are followed by other low values and "high" values (i.e., those fluctuating around a mean of 6) are followed by other high values. For a formal demonstration, note that as a_1 approaches unity, (4.30) approaches a random walk plus drift. We know that the solution to the random walk plus drift model includes a deterministic trend, that is,

Random walk with a structural break:

2

$$y_t = y_{t-1} + \varepsilon_t + D_p$$

$$D_p(51) = 4 \text{ and } 0 \text{ otherwise}$$

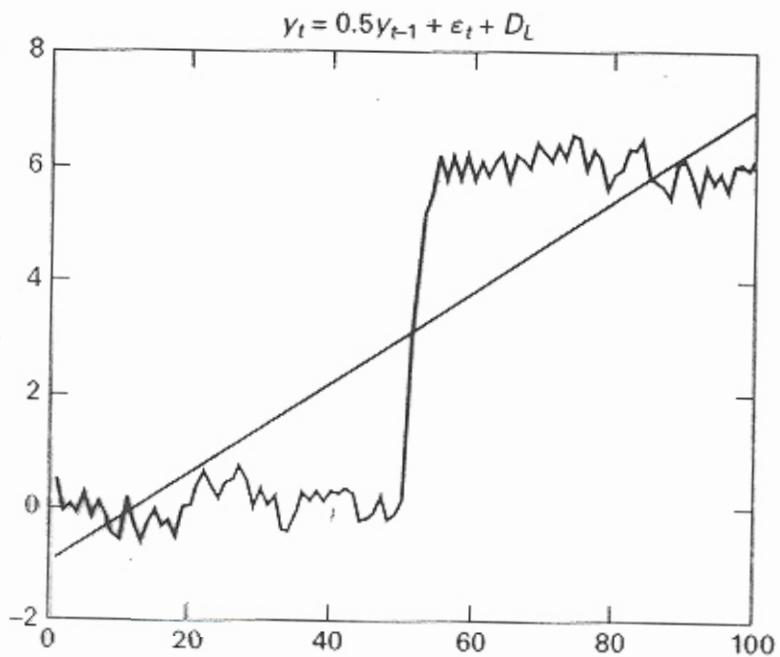
Test procedure.

H_0 : one-time jump in the level of a unit root process

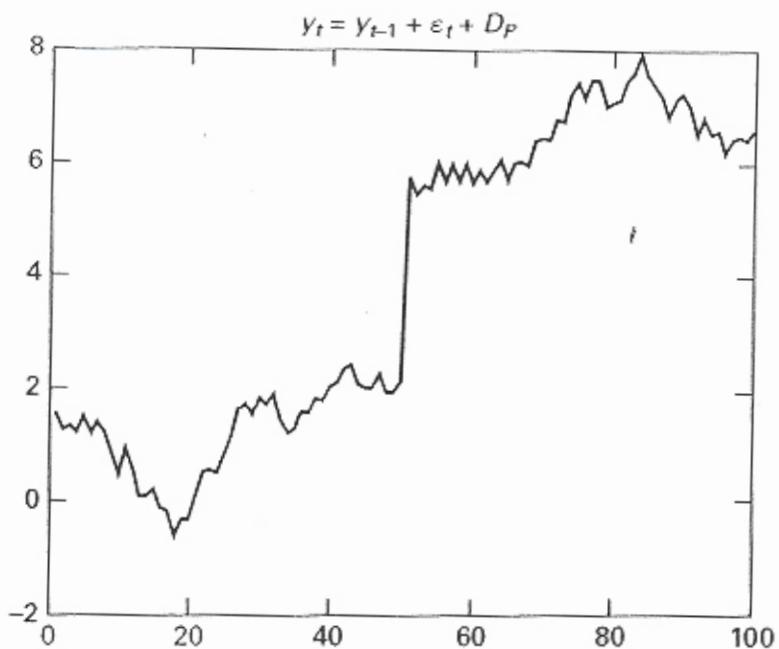
H_A : a one-time change in the intercept of a trend stationary process.

$$H_0: y_t = y_0 + y_{t-1} + M_1 D_p + \varepsilon_t$$

$$H_A: y_t = a_0 + a_2 t + M_2 D_L + \varepsilon_t$$



$$D_L = \begin{cases} 1 & \text{if } t > T \\ 0 & \text{otherwise} \end{cases}$$



$$D_P = \begin{cases} 1 & \text{if } t = T+1 \\ 0 & \text{otherwise} \end{cases}$$

FIGURE 4.10 Two Models of Structural Change

$$H_0: y_t = a_0 + y_{t-1} + \mu_1 D_p + \varepsilon_t$$

$$H_A: y_t = a_0 + a_2 t + \mu_2 D_L + \varepsilon_t$$

STEP 1: Detrend the data by estimating the model under the alternative hypothesis and calling the residuals \hat{y}_t . Hence, each value of \hat{y}_t is the residual from the regression $y_t = a_0 + a_2 t + \mu_2 D_L + \hat{y}_t$.

STEP 2: Estimate the regression:

$$\hat{y}_t = a_1 \hat{y}_{t-1} + \varepsilon_t$$

Under the null hypothesis of a unit root, the theoretical value of a_1 is unity. Perron (1989) shows that when the residuals are identically and independently distributed, the distribution of a_1 depends on the proportion of observations occurring prior to the break. Denote this proportion by $\lambda = \tau/T$ where T = total number of observations.

STEP 3: Perform diagnostic checks to determine if the residuals from Step 2 are serially uncorrelated. If there is serial correlation, use the augmented form of the regression:

$$\hat{y}_t = a_1 \hat{y}_{t-1} + \sum_{i=1}^k \beta_i \Delta \hat{y}_{t-i} + \varepsilon_t$$

STEP 4: Calculate the t -statistic for the null hypothesis $a_1 = 1$. This statistic can be compared to the critical values calculated by Perron. Perron generated 5,000 series according to H_1 using values of λ ranging from 0 to 1 by increments of 0.1. For each value of λ , he estimated the regressions $\hat{y}_t = a_1 \hat{y}_{t-1} + \varepsilon_t$ and calculated the sample distribution of a_1 . Naturally the critical values are identical to the Dickey-Fuller statistics when $\lambda = 0$ and $\lambda = 1$; in effect,

there is no structural change unless $0 < \lambda < 1$. The maximum difference between the two statistics occurs when $\lambda = 0.5$. For $\lambda = 0.5$, the critical value of the t -statistic at the 5 percent level of significance is -3.76 (which is larger in absolute than the corresponding Dickey-Fuller statistic -3.41). If you find a t -statistic greater than the critical value calculated by Perron, it is possible to reject the null hypothesis of a unit root.

SPURIOUS REGRESSION

Greniger, Newbold (1979)

$$Y_t = a_0 + a_1 Z_t + \epsilon_t$$

- has high R^2
 - t statistics are significant
 - no economic meaning
- ⇒ looks good because OLS INCONSISTENT
ALL INFERENCE IS WRONG.

- IF y_t, z_t have stochastic trends, $\{\epsilon_t\}$ may be nonstationary:
 - i) IF $\{\epsilon_t\}$ has a stochastic trend - each deviation is permanent since the error in any period t never decays.

$$y_t = y_{t-1} + \epsilon_{yt}$$

$$z_t = z_{t-1} + \epsilon_{zt}$$

$$e_t = y_t - a_1 z_t$$

$$e_t = \sum_{i=1}^+ \epsilon_{yi} - a_1 \sum_{i=1}^+ \epsilon_{zi}$$

- the $\text{var}(e_t) \rightarrow \infty$ as $t \rightarrow \infty$

- $E e_{t+1} = e_t$

- strongly correlated

A. IF $\{y_t\}$ $\{z_t\}$ INTEGRATED OF DIFFERENT ORDERS \Rightarrow regression is meaningless.

B. $\{y_t\}$ $\{z_t\}$ INTEGRATED OF THE SAME ORDER AND RETURNS HAVE A STOCHASTIC TREND
 \Rightarrow estimate $\Delta y_t = a_1 \Delta z_t + \Delta e_t$

c. $\{y_t\}$ $\{z_t\}$ INTEGRATED OF THE SAME ORDER
AND $\{\epsilon_t\}$ STATIONARY \Rightarrow COINTEGRATION

$$y_t = M_t + \varepsilon_{y_t}$$

$$z_t = M_t + \varepsilon_{z_t}$$

$$M_t = M_{t-1} + \varepsilon_M$$

$$y_t - z_t = \varepsilon_{y_t} - \varepsilon_{z_t} \text{ IS STATIONARY}$$